

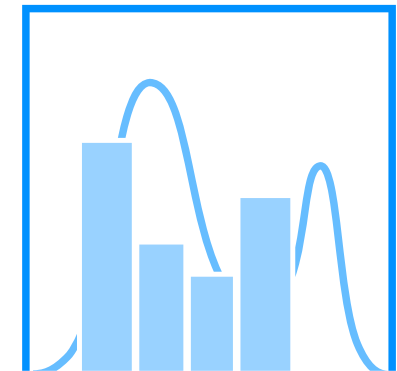
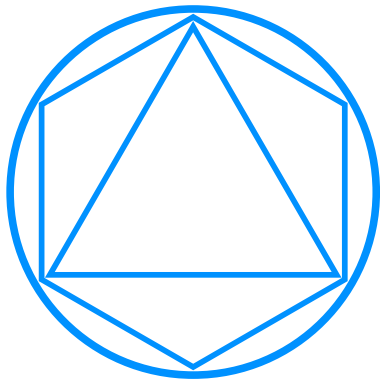
# Multivariate Lévy driven Stochastic Volatility Models

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Parts based on joint work with O. E. Barndorff-Nielsen and Ch. Pigorsch

# Outline of this talk

- Motivation from finance and the univariate model
- Matrix subordinators
- Positive semi-definite Ornstein-Uhlenbeck type processes (based on Barndorff-Nielsen & St., 2007; Pigorsch & St., 2008a)
- **Multivariate Ornstein-Uhlenbeck type stochastic volatility model** (based on Pigorsch & St., 2008b)
- **Multivariate COGARCH(1,1)** (based on St., 2008)

# Stylized Facts of Financial Return Data

- non-constant, **stochastic volatility**
- volatility exhibits **jumps**
- asymmetric and heavily tailed marginal distributions
- clusters of extremes
- **log returns exhibit marked dependence, but have vanishing autocorrelations**  
(squared returns, for instance, have non-zero autocorrelation)

**Stochastic Volatility Models** are used to cover these stylized facts.

# Univariate BNS Model I

- **Logarithmic stock price** process  $(Y_t)_{t \in \mathbb{R}^+}$ :

$$dY_t = (\mu + \beta \sigma_{t-}) dt + \sigma_{t-}^{1/2} dW_t$$

with parameters  $\mu, \beta \in \mathbb{R}$  and  $(W_t)_{t \in \mathbb{R}^+}$  being standard Brownian motion.

- Ornstein-Uhlenbeck-type **volatility** process  $(\sigma_t)_{t \in \mathbb{R}^+}$ :

$$d\sigma_t = -\lambda \sigma_{t-} dt + dL_t, \quad \sigma_0 > 0$$

with parameter  $\lambda > 0$  and  $(L_t)_{t \in \mathbb{R}^+}$  being a **Lévy subordinator**.

## Univariate BNS Model II

- Usually  $E(\max(\log |L_1|, 0)) < \infty$  and  $\sigma$  is chosen as the unique stationary solution to  $d\sigma_t = -\lambda\sigma_t dt + dL_t$  given by

$$\sigma_t = \int_{-\infty}^t e^{-\lambda(t-s)} dL_s.$$

- Closed form expression for the integrated volatility

$$\int_0^t \sigma_s ds = \frac{1}{\lambda} (L_t - \sigma_t + \sigma_0).$$

Derivative Pricing via Laplace transforms possible.

# The Need for Multivariate Models

Multivariate models are needed

- to study **comovements and spill over effects** between several assets.
- for optimal **portfolio selection and risk management** at a portfolio level.
- to price **derivatives on multiple assets**.

**Desire:**

Multivariate models that are **flexible, realistic and analytically tractable**.

# Some Matrix Notation

- $M_d(\mathbb{R})$ : the real  $d \times d$  matrices.
- $\mathbb{S}_d$ : the real symmetric  $d \times d$  matrices.
- $\mathbb{S}_d^+$ : the positive-semidefinite  $d \times d$  matrices (covariance matrices) (a closed cone).
- $\mathbb{S}_d^{++}$ : the positive-definite  $d \times d$  matrices (an open cone).
- $A^{1/2}$ : for  $A \in \mathbb{S}_d^+$  the unique positive-semidefinite square root (functional calculus).

# Matrix Subordinators

- **Definition:**

An  $\mathbb{S}_d$ -valued Lévy process  $L$  is said to be a *matrix subordinator*, if  $L_t - L_s \in \mathbb{S}_d^+$  for all  $s, t \in \mathbb{R}^+$  with  $t > s$ .

(Barndorff-Nielsen and Pérez-Abreu (2008)).

- The paths are  $\mathbb{S}_d^+$ -increasing and of finite variation.
- The characteristic function  $\mu_{L_t}$  of  $L_t$  for  $t \in \mathbb{R}^+$  is given by

$$\mu_{L_t}(Z) = \exp \left( t \left( i \operatorname{tr}(\gamma_L Z) + \int_{\mathbb{S}_d^+ \setminus \{0\}} \left( e^{i \operatorname{tr}(XZ)} - 1 \right) \nu_L(dX) \right) \right), \quad Z \in \mathbb{S}_d,$$

where  $\gamma_L$  is the drift and  $\nu_L$  the Lévy measure.



# Examples of Matrix Subordinators

- **Analogues of univariate subordinators** can be defined via the characteristic functions: e.g. (tempered) stable, Gamma or IG matrix subordinators
- **Diagonal matrix subordinators**, i.e. off-diagonal elements zero, diagonal elements univariate subordinators
- Discontinuous part of the **Quadratic (Co-)Variation** process of any  $d$ -dimensional Lévy process  $\tilde{L}$ :

$$[\tilde{L}, \tilde{L}]_t^{\text{d}} = \sum_{s \leq t} \Delta \tilde{L}_s (\Delta \tilde{L}_s)^T$$

# Linear Operators Preserving Positive-Semidefiniteness

**Proposition** Let  $\mathbf{A} : \mathbb{S}_d \rightarrow \mathbb{S}_d$  be a linear operator. Then  $e^{\mathbf{A}t}(\mathbb{S}_d^+) = \mathbb{S}_d^+$  for all  $t \in \mathbb{R}$ , if and only if  $\mathbf{A}$  is representable as  $X \mapsto AX + XA^T$  for some  $A \in M_d(\mathbb{R})$ . □

One has  $e^{\mathbf{A}t}X = e^{At}Xe^{A^T t}$  for all  $X \in \mathbb{S}_d$ .

In the above setting  $\sigma(\mathbf{A}) = \sigma(A) + \sigma(A)$ . Hence,  $\mathbf{A}$  has only eigenvalues of strictly negative real part, if and only if this is the case for  $A$ .

## Positive-semidefinite OU-type Processes

**Theorem** Let  $(L_t)_{t \in \mathbb{R}}$  be a **matrix subordinator** with  $E(\max(\log \|L_1\|, 0)) < \infty$  and  $A \in M_d(\mathbb{R})$  such that  $\sigma(A) \subset (-\infty, 0) + i\mathbb{R}$ .

Then the stochastic differential equation of Ornstein-Uhlenbeck-type

$$d\Sigma_t = (A\Sigma_{t-} + \Sigma_{t-}A^T)dt + dL_t$$

has a **unique stationary solution**

$$\Sigma_t = \int_{-\infty}^t e^{A(t-s)} dL_s e^{A^T(t-s)}$$

or, in vector representation,  $\text{vec}(\Sigma_t) = \int_{-\infty}^t e^{(I_d \otimes A + A \otimes I_d)(t-s)} d\text{vec}(L_s)$ .

Moreover,  $\Sigma_t \in \mathbb{S}_d^+$  for all  $t \in \mathbb{R}$ . □

## Stationary Distribution

**Theorem** Let  $\gamma_L$  be the drift of the driving matrix subordinator  $L$  and  $\nu_L$  its Lévy measure.

The stationary distribution of the Ornstein-Uhlenbeck process  $\Sigma$  is infinitely divisible (even operator self-decomposable) with characteristic function

$$\hat{\mu}_\Sigma(Z) = \exp \left( i \operatorname{tr}(\gamma_\Sigma Z) + \int_{\mathbb{S}_d^+ \setminus \{0\}} (e^{i \operatorname{tr}(YZ)} - 1) \nu_\Sigma(dY) \right), \quad Z \in \mathbb{S}_d,$$

where

$$\gamma_\Sigma = -\mathbf{A}^{-1} \gamma_L \quad \text{and} \quad \nu_\Sigma(E) = \int_0^\infty \int_{\mathbb{S}_d^+ \setminus \{0\}} I_E(e^{As} x e^{A^T s}) \nu_L(dx) ds$$

for all Borel sets  $E$  in  $\mathbb{S}_d^+ \setminus \{0\}$ .

$\mathbf{A}^{-1}$  is the inverse of the linear operator  $\mathbf{A} : \mathbb{S}_d(\mathbb{R}) \rightarrow \mathbb{S}_d(\mathbb{R})$ ,  $X \mapsto AX + XA^T$  which can be represented as  $\operatorname{vec}^{-1} \circ ((I_d \otimes A) + (A \otimes I_d))^{-1} \circ \operatorname{vec}$ .  $\square$

## Strict Positive-definiteness

**Proposition** If  $\gamma_L \in \mathbb{S}_d^{++}$  or  $\nu_L(\mathbb{S}_d^{++}) > 0$ , then the stationary distribution  $P_\Sigma$  of  $\Sigma$  is concentrated on  $\mathbb{S}_d^{++}$ , i.e.  $P_\Sigma(\mathbb{S}_d^{++}) = 1$ .  $\square$

**Theorem** Let  $\tilde{L}$  be a Lévy process in  $\mathbb{R}^d$  with Lévy measure  $\nu_{\tilde{L}} \neq 0$  and assume that  $\nu_{\tilde{L}}$  is absolutely continuous (with respect to the Lebesgue measure on  $\mathbb{R}^d$ ).

Then the stationary distribution of the Ornstein-Uhlenbeck type process  $\Sigma_t$  driven by the discontinuous part of the quadratic variation  $[\tilde{L}, \tilde{L}]_t^\partial$  is absolutely continuous with respect to the Lebesgue measure. Moreover, the stationary distribution  $P_\Sigma$  of  $\Sigma_t$  is concentrated on  $\mathbb{S}_d^{++}$ , i.e.  $P_\Sigma(\mathbb{S}_d^{++}) = 1$ .  $\square$

## Marginal Dynamics

Assume that  $A$  is real diagonalisable and let  $U \in GL_d(\mathbb{R})$  be such that  $UAU^{-1} =: D$  is diagonal.

- $M_t := UL_tU^T$  is again a matrix subordinator.
- $(U\Sigma_tU^T)_{ij} = \left( \int_{-\infty}^t e^{D(t-s)} d(UL_sU^T) e^{D(t-s)} \right)_{ij} = \int_{-\infty}^t e^{(\lambda_i + \lambda_j)(t-s)} dM_{ij,s}$ .
- Hence, the individual components of  $U\Sigma_tU^T$  are stationary one-dimensional Ornstein-Uhlenbeck type processes with associated SDE  $d(U\Sigma_tU^T)_{ij} = (\lambda_i + \lambda_j)(U\Sigma_tU^T)_{ij}dt + dM_{ij,t}$ .

$M_{ii}$  for  $1 \leq i \leq d$  are necessarily subordinators and  $(U\Sigma_tU^T)_{ii}$  have to be positive OU type processes.

- The individual components  $\Sigma_{ij,t}$  of  $\Sigma_t$  are superpositions of (at most  $d^2$ ) univariate OU type processes. The individual OU processes superimposed are in general not independent.

## Second Order Structure

**Theorem** Assume that the driving Lévy process is square-integrable. Then the second order moment structure is given by

$$\begin{aligned}
 E(\Sigma_t) &= \gamma_\Sigma - \mathbf{A}^{-1} \int_{\mathbb{S}_d^+ \setminus \{0\}} y \nu(dy) = -\mathbf{A}^{-1} E(L_1) \\
 \text{var}(\text{vec}(\Sigma_t)) &= -\mathcal{A}^{-1} \text{var}(\text{vec}(L_1)) \\
 \text{cov}(\text{vec}(\Sigma_{t+h}), \text{vec}(\Sigma_t)) &= e^{(\mathbf{A} \otimes I_d + I_d \otimes \mathbf{A})h} \text{var}(\text{vec}(\Sigma_t)),
 \end{aligned}$$

where  $t \in \mathbb{R}$  and  $h \in \mathbb{R}^+$ ,  $\mathbf{A} : M_d(\mathbb{R}) \rightarrow M_d(\mathbb{R})$ ,  $X \mapsto AX + XA^T$  and  $\mathcal{A} : M_{d^2}(\mathbb{R}) \rightarrow M_{d^2}(\mathbb{R})$ ,  $X \mapsto (\mathbf{A} \otimes I_d + I_d \otimes \mathbf{A})X + X(\mathbf{A}^T \otimes I_d + I_d \otimes \mathbf{A}^T)$ .  $\square$

The individual components of the autocovariance matrix do not have to decay exponentially, but may exhibit exponentially damped sinusoidal behaviour.

# The Integrated Volatility

**Theorem** The integrated Ornstein-Uhlenbeck process  $\Sigma_t^+$  is given by

$$\Sigma_t^+ := \int_0^t \Sigma_t dt = \mathbf{A}^{-1} (\Sigma_t - \Sigma_0 - L_t)$$

for  $t \in \mathbb{R}^+$ .





# Multivariate OU type Stochastic Volatility Model

$d$ -dimensional logarithmic stock price process  $(Y_t)_{t \in \mathbb{R}}$ :

$$dY_t = (\mu + \Sigma_{t-}\beta) dt + \Sigma_{t-}^{1/2} dW_t$$

with

- $(W_t)_{t \in \mathbb{R}^+}$  being  $d$ -dimensional standard Brownian motion,
- $\mu, \beta \in \mathbb{R}^d$  and
- $(\Sigma_t)_{t \in \mathbb{R}^+}$  being a stationary  $\mathbb{S}_d^+$ -valued Ornstein-Uhlenbeck type process.

$\implies$  Natural analogue of the univariate model

# The Conditional Fourier Transform

Assume the driving matrix subordinator  $L$  has characteristic exponent  $\psi_L$ , i.e.  $E(e^{i\text{tr}(L_t z)}) = e^{t\psi_L(z)}$  for all  $z \in M_d(\mathbb{R}) + i\mathbb{S}_d^+$ . Let  $(Y_0, \Sigma_0) \in \mathbb{R}^d \times \mathbb{S}_d^+$  be the initial values. Then we have for every  $t \in \mathbb{R}^+$  and  $(y, z) \in \mathbb{R}^d \times M_d(\mathbb{R})$

$$\begin{aligned}
 E \left( e^{i(Y_t^T y + \text{tr}(\Sigma_t z))} \middle| \Sigma_0, Y_0 \right) &= \exp \left\{ i(Y_0 + \mu t)^T y + i \text{tr} \left( \Sigma_0 e^{A^T t} z e^{At} \right) \right. \\
 &+ i \text{tr} \left( \Sigma_0 e^{A^T t} \left[ \mathbf{A}^{-*} \left( y \beta^T + \frac{i}{2} y y^T \right) \right] e^{At} - \Sigma_0 \left[ \mathbf{A}^{-*} \left( y \beta^T + \frac{i}{2} y y^T \right) \right] \right) \\
 &\left. + \int_0^t \psi_L \left( e^{A^T s} z e^{As} + e^{A^T s} \left[ \mathbf{A}^{-*} \left( y \beta^T + \frac{i}{2} y y^T \right) \right] e^{As} - \mathbf{A}^{-*} \left( y \beta^T + \frac{i}{2} y y^T \right) \right) ds \right\}
 \end{aligned}$$

with  $\mathbf{A}^{-*}$  denoting the inverse of the adjoint of  $\mathbf{A}$ , i.e.  $\mathbf{A}^{-*}$  is the inverse of the linear operator  $\mathbf{A}^*$  given by  $X \mapsto A^T X + X A$ .

# The Logarithmic Returns

Let  $\Delta > 0$  (grid size). Define for  $n \in \mathbb{N}$ :

- log-returns over periods  $[(n-1)\Delta, n\Delta]$  of length  $\Delta$ :

$$\mathbf{Y}_n = Y_{n\Delta} - Y_{(n-1)\Delta} = \int_{(n-1)\Delta}^{n\Delta} (\mu + \Sigma_t \beta) dt + \int_{(n-1)\Delta}^{n\Delta} \Sigma_t^{1/2} dW_t.$$

- Integrated volatility over  $[(n-1)\Delta, n\Delta]$ :

$$\Sigma_n := \int_{(n-1)\Delta}^{n\Delta} \Sigma_t dt.$$

It holds that

$$\mathbf{Y}_n | \Sigma_n \sim N_d(\mu\Delta + \Sigma_n \beta, \Sigma_n)$$

with  $N_d$  denoting the  $d$ -dimensional normal distribution.

## Second Order Structure of $\Sigma_n$

Assume henceforth  $E(\|L_1\|^2) < \infty$ .

$$\begin{aligned}
 E(\Sigma_n) &= \Delta E(\Sigma_0) = -\Delta \mathbf{A}^{-1} E(L_1) \\
 \text{var}(\text{vec}(\Sigma_n)) &= r^{++}(\Delta) + (r^{++}(\Delta))^T \\
 r^{++}(t) &= (\mathcal{A}^{-2} (e^{\mathcal{A}t} - I_{d^2}) - \mathcal{A}^{-1}t) \text{var}(\text{vec}(\Sigma_0)) \\
 &= -(\mathcal{A}^{-2} (e^{\mathcal{A}t} - I_{d^2}) - \mathcal{A}^{-1}t) \mathcal{A}^{-1} \text{var}(\text{vec}(L_1)) \\
 \text{acov}_\Sigma(h) &= e^{\mathcal{A}\Delta(h-1)} \mathcal{A}^{-2} (I_{d^2} - e^{\mathcal{A}\Delta})^2 \text{var}(\text{vec}(\Sigma_0)) \\
 &= -e^{\mathcal{A}\Delta(h-1)} \mathcal{A}^{-2} (I_{d^2} - e^{\mathcal{A}\Delta})^2 \mathcal{A}^{-1} \text{var}(\text{vec}(L_1)), \quad h \in \mathbb{N}.
 \end{aligned}$$

where  $\mathcal{A} = A \otimes I_d + I_d \otimes A$  and  $\mathcal{A} : M_{d^2}(\mathbb{R}) \rightarrow M_{d^2}(\mathbb{R}), X \mapsto \mathcal{A}X + X\mathcal{A}^T$ .

$\implies \text{vec}(\Sigma_n)$  is a **causal ARMA(1,1) process** with AR parameter  $e^{\mathcal{A}\Delta}$ .

## Second Order Structure of $\mathbf{Y}_n$ and $\mathbf{Y}_n \mathbf{Y}_n^T$

$$E(\mathbf{Y}_n) = (\mu + E(\Sigma_0)\beta)\Delta$$

$$\text{var}(\mathbf{Y}_n) = E(\Sigma_0)\Delta + (\beta^T \otimes I_d)\text{var}(\text{vec}(\Sigma_n))(\beta \otimes I_d)$$

$$\text{acov}_{\mathbf{Y}}(h) = (\beta^T \otimes I_d)\text{acov}_{\Sigma}(h)(\beta \otimes I_d), \quad h \in \mathbb{N}$$

Assume  $\mu = \beta = 0$ . Then:

$$E(\mathbf{Y}_n \mathbf{Y}_n^T) = E(\Sigma_0)\Delta$$

$$\begin{aligned} \text{var}(\text{vec}(\mathbf{Y}_n \mathbf{Y}_n^T)) &= (I_{d^2} + \mathbf{Q} + \mathbf{PQ}) \text{var}(\text{vec}(\Sigma_n)) \\ &\quad + (I_{d^2} + \mathbf{P})(E(\Sigma_0) \otimes E(\Sigma_0)) \Delta^2 \end{aligned}$$

$$\text{acov}_{\mathbf{Y}\mathbf{Y}^T}(h) = \text{acov}_{\Sigma}(h) \text{ for } h \in \mathbb{N}$$

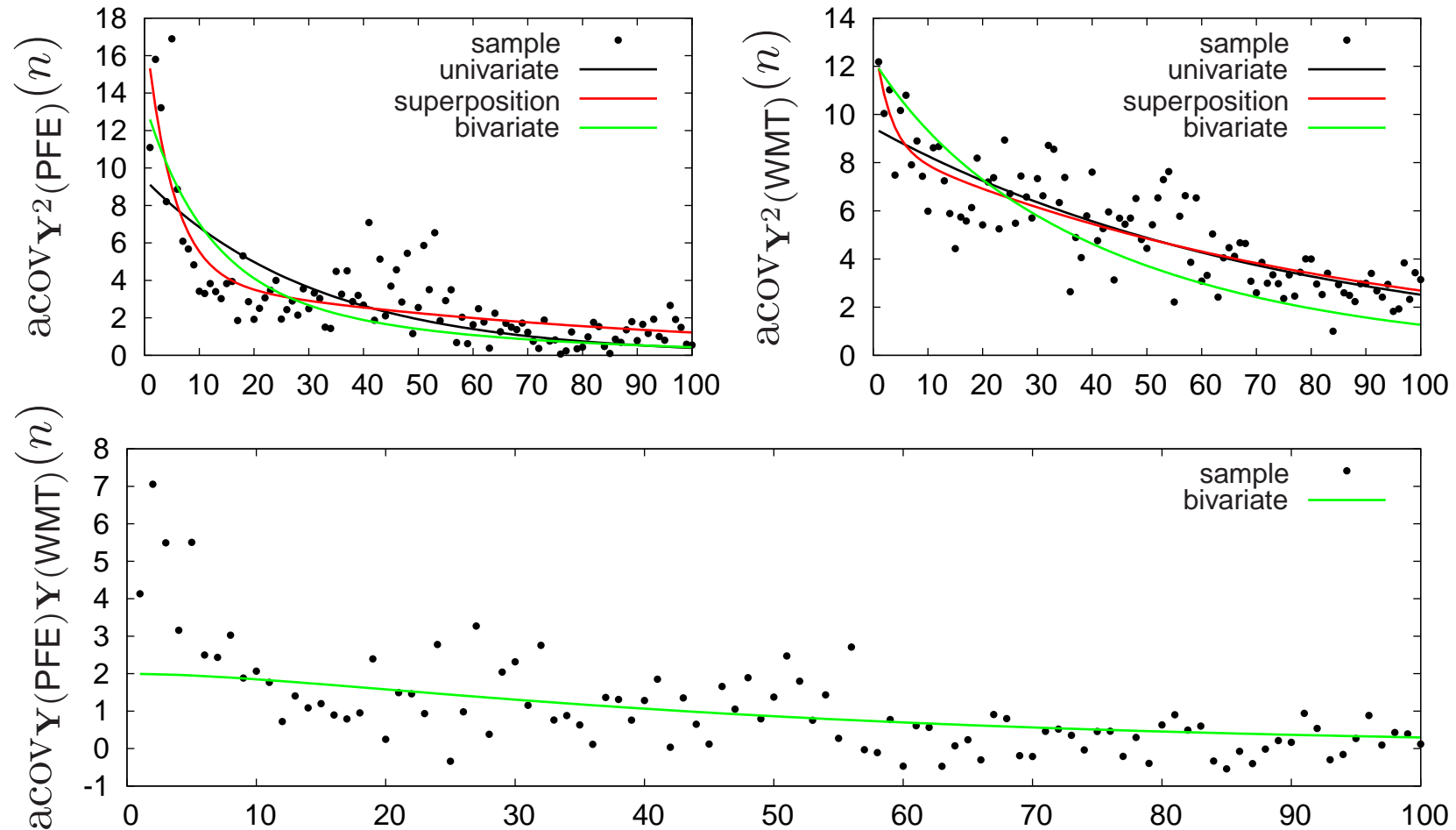
where  $\mathbf{P}$  and  $\mathbf{Q}$  are linear operators on  $M_{d^2}(\mathbb{R})$  rearranging the entries.

$\implies \text{vec}(\mathbf{Y}_n \mathbf{Y}_n^T)$  is a **causal ARMA(1,1) process** with AR parameter  $e^{\mathcal{A}\Delta}$ .

# Moment Estimators

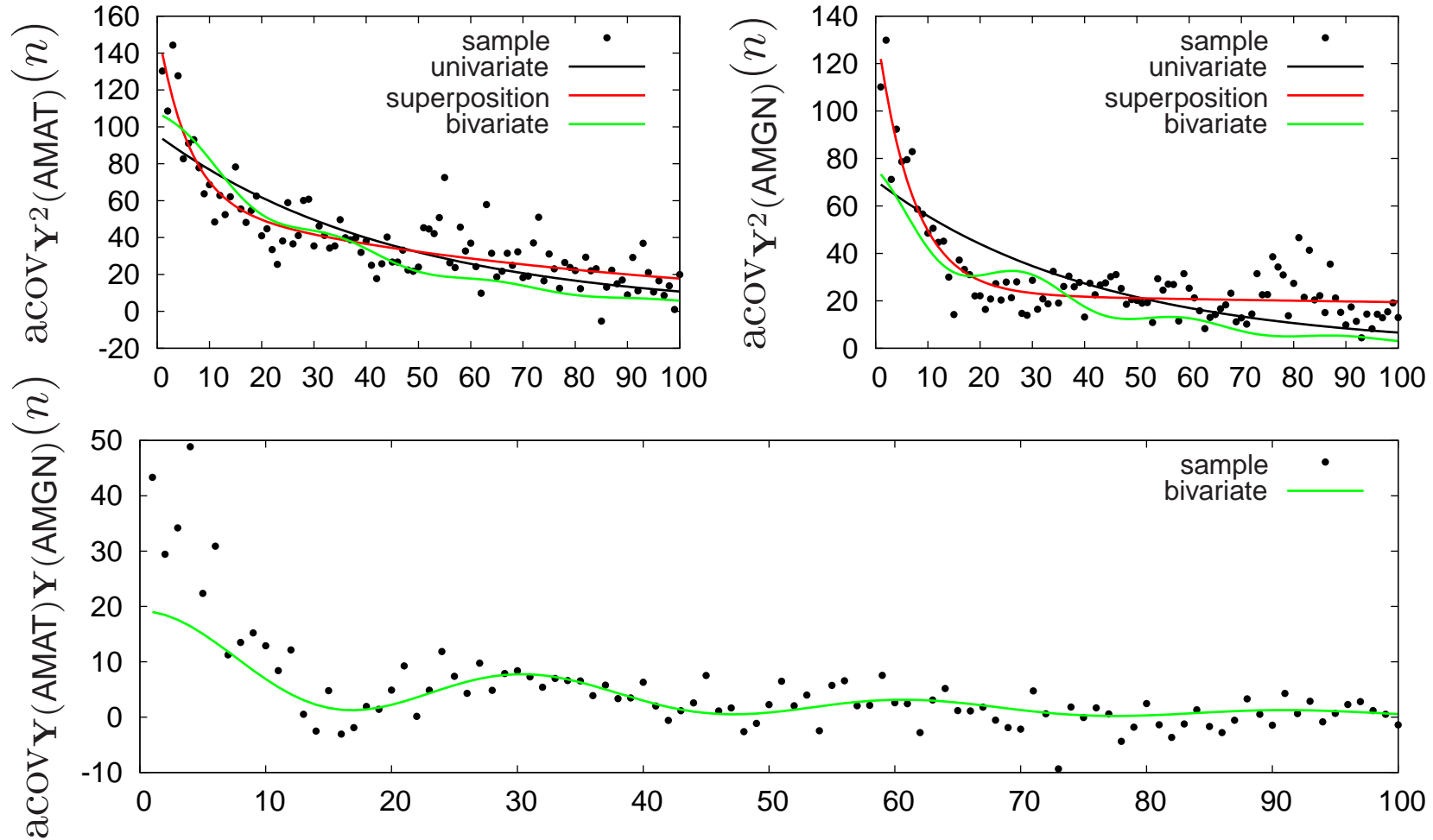
- Assume  $\mu = \beta = 0$
- $E(L_1)$ ,  $\text{var}(\text{vec}(L_1))$  and  $A$  can be estimated from the empirically observed  $E(\mathbf{Y}_n \mathbf{Y}_n^T)$ ,  $\text{acov}_{\mathbf{Y}\mathbf{Y}^T}(1)$  and  $\text{acov}_{\mathbf{Y}\mathbf{Y}^T}(2)$ .
- They are identified provided one assumes that  $e^{\mathbf{A}_{\text{vech}}\Delta}$  has a unique real logarithm and  $\text{var}(\text{vech}(\Sigma_0))$  is invertible.
- In practice one uses more lags of the autocovariance function and GMM estimation.
- **The log-returns  $\mathbf{Y}$  are strongly mixing.** Thus the estimators are under appropriate technical conditions consistent and asymptotically normal.

# Empirical Illustration I



Empirical and estimated autocovariance functions: PFE and WMT

# Empirical Illustration II



Empirical and estimated autocovariance functions: AMAT and AMGN



## Extensions

Even more flexibility (and **long memory**) by considering superpositions of independent multivariate positive semi-definite OU type processes for  $\Sigma$ .

Possibilities:

- Superposition of finitely many OU type processes: Straightforward and (almost) all results easily extendible.
- Superposition of countably many OU type processes and convergence in  $L^2$ .
- Use of a  $\mathbb{S}_d^+$ -valued Lévy basis  $\Lambda$  on  $\mathbb{R} \times M_d^-(\mathbb{R})$  with  $M_d^-(\mathbb{R}) := \{X \in M_d(\mathbb{R}) : \sigma(X) \subset (-\infty, 0) + i\mathbb{R}\}$ :

$$\Sigma_t = \int_{-\infty}^t \int_{M_d^-(\mathbb{R})} e^{A(t-s)} \Lambda(ds, dA) e^{A^T(t-s)}$$

## Univariate BNS and COGARCH model

- The **Ornstein-Uhlenbeck type stochastic volatility** model (BNS model):

$$dY_t = \sqrt{\sigma_{t-}} dW_t$$

$$d\sigma_t = -\lambda\sigma_{t-}dt + dL_t$$

with  $\lambda > 0$ ,  $W$  standard Brownian motion and  $L$  a subordinator.

- The **COGARCH(1,1)** model (Klüppelberg, Lindner, Maller (2004)):

$$dY_t = \sqrt{\sigma_{t-}} dL_t$$

$$\sigma_t = c + v_t, \quad dv_t = -\alpha v_{t-}dt + \beta\sigma_{t-}d[L, L]_t^\diamond$$

with  $\alpha, \beta, c > 0$ ,  $L$  a Lévy process and  $[L, L]_t^\diamond = \sum_{0 < s \leq t} (\Delta L_s)^2$ .

## Multivariate COGARCH(1,1) – Definition

**Definition** Let  $L$  be a  $d$ -dimensional Lévy process and  $A, B \in M_d(\mathbb{R})$ ,  $C \in \mathbb{S}_d^+$  and set  $[L, L]_t^\diamond := \sum_{0 < s \leq t} \Delta L_s (\Delta L_s)^T$ . Then the process  $Y = (Y_t)_{t \in \mathbb{R}^+}$  solving

$$dY_t = \Sigma_{t-}^{1/2} dL_t, \quad \Sigma_t = C + V_t, \quad (1)$$

$$dV_t = (AV_{t-} + V_{t-}A^T)dt + B\Sigma_{t-}^{1/2} d[L, L]_t^\diamond \Sigma_{t-}^{1/2} B^T \quad (2)$$

with initial values  $Y_0 = 0$  in  $\mathbb{R}^d$  and  $V_0$  in  $\mathbb{S}_d^+$  is called a *multivariate COGARCH(1,1) process*.

The process  $V = (V_t)_{t \in \mathbb{R}^+}$  (or  $\Sigma$ ) with paths in  $\mathbb{S}_d^+$  is referred to as a *multivariate COGARCH(1,1) volatility process*.  $\square$

Agrees with the definition of the COGARCH(1,1) for  $d = 1$  and inherits many of the properties of multivariate GARCH(1,1).

# Multivariate COGARCH(1,1) – Equivalent Definitions

- One can directly define  $\Sigma$  via the SDE

$$d\Sigma_t = (A(\Sigma_{t-} - C) + (\Sigma_{t-} - C)A^T)dt + B\Sigma_{t-}^{1/2}d[L, L]_t^{\circ}\Sigma_{t-}^{1/2}B^T$$

which shows that  $\Sigma$  has a mean reverting structure (provided  $\sigma(A) \subset (-\infty, 0) + i\mathbb{R}$ ) with “mean”  $C$ .

- The volatility process  $V$  (or  $\Sigma$ ) is of finite variation and  $V$  satisfies for all  $t \in \mathbb{R}^+$

$$V_t = e^{At}V_0e^{A^T t} + \int_0^t e^{A(t-s)}B\Sigma_{s-}^{1/2}d[L, L]_s^{\circ}\Sigma_{s-}^{1/2}B^T e^{A^T(t-s)}.$$

# Markovian Properties and Stationarity

Provided  $C \in \mathbb{S}_d^{++}$ ,  $(Y, V)$  and  $V$  alone are **temporally homogeneous strong Markov processes** on  $\mathbb{R}^d \times \mathbb{S}_d^+$  and  $\mathbb{S}_d^+$ , respectively. Moreover, both have the **weak Feller** property.

**Theorem 1.** *Assume:*

- $C \in \mathbb{S}_d^{++}$ ,  $A \in M_d(\mathbb{R})$  is *diagonalisable* with  $S \in GL_d(\mathbb{C})$  such that  $S^{-1}AS$  is diagonal,
- the Lévy measure  $\nu_L$  of  $L$  satisfies

$$\int_{\mathbb{R}^d} \log \left( 1 + \alpha_1 \| (S^{-1} \otimes S^{-1}) \text{vec}(yy^T) \|_2 \right) \nu_L(dy) < -2 \max(\Re(\sigma(A))),$$

$$\text{where } \alpha_1 := \|S\|_2^2 \|S^{-1}\|_2^2 K_{2,A} \| (S^{-1}BS) \otimes (S^{-1}BS) \|_2,$$

$$K_{2,A} := \max_{X \in \mathbb{S}_d^+, \|X\|_2=1} \left( \frac{\|X\|_2}{\| (S^{-1} \otimes S^{-1}) \text{vec}(X) \|_2} \right).$$

Then *there exists a stationary distribution  $\mu$  for the multivariate COGARCH(1,1) volatility process  $V$  having the following property:*

*If*

$$\int_{\mathbb{R}^d} \left( (1 + \alpha_1 \| (S^{-1} \otimes S^{-1}) \text{vec}(yy^T) \|_2)^k - 1 \right) \nu_L(dy) < -2k \max(\Re(\sigma(A)))$$

*for some  $k \in \mathbb{N}$ , then*

$$\int_{\mathbb{S}_d^+} \|x\|^k \mu(dx) < \infty,$$

*i.e. the  $k$ -th moment of  $\mu$  is finite.*

The stationary distribution is not known to be unique or to be a limiting distribution.

## Second Order Properties

Assume:

- The driving Lévy process  $L$  has finite fourth moments and  $\nu_L$  satisfies

$$\int_{\mathbb{R}^d} xx^T \nu_L(dx) = \sigma_L I_d,$$

$$\int_{\mathbb{R}^d} \text{vec}(xx^T) \text{vec}(xx^T)^T \nu_L(dx) = \rho_L (I_{d^2} + K_d + \text{vec}(I_d) \text{vec}(I_d)^T)$$

for some  $\sigma_L, \rho_L \in \mathbb{R}^+$  and with  $K_d$  being the commutation matrix,

- $\sigma(A), \sigma(\mathcal{A}), \sigma(\mathcal{C}) \subset (-\infty, 0) + i\mathbb{R}$  with

$$\mathcal{A} = A \otimes I_d + I_d \otimes A + \sigma_L B \otimes B,$$

$$\mathcal{C} := \mathcal{A} \otimes I_{d^2} + I_{d^2} \otimes \mathcal{A} + \sigma_L ((B \otimes B) \otimes I_{d^2} + I_{d^2} \otimes (B \otimes B)) + \mathcal{BR},$$

$$\mathcal{B} = (B \otimes B) \otimes (B \otimes B), \quad \mathcal{R} = \rho_L (\mathcal{Q} + \mathcal{K}_d \mathcal{Q} + I_{d^4}),$$

where  $\mathcal{K}_d$  and  $\mathcal{Q}$  are certain permutation matrices.

- $V_0$  has finite second moments.

Then  $V$  is asymptotically second order stationary with

- mean

$$E(\text{vec}(V_\infty)) = -\sigma_L \mathcal{A}^{-1} (B \otimes B) \text{vec}(C),$$

- autocovariance function

$$\text{acov}_{\text{vec}(V_\infty)}(h) = e^{\mathcal{A}h} \text{var}(\text{vec}(V_\infty))$$

for  $h \in \mathbb{R}^+$

- and variance

$$\begin{aligned} \text{vec}(\text{var}(\text{vec}(V_\infty))) &= -\mathcal{L}^{-1} [(\sigma_L^2 \mathcal{L}(\mathcal{A}^{-1} \otimes \mathcal{A}^{-1}) \mathcal{B} + \mathcal{B}\mathcal{R}) (\text{vec}(C) \otimes \text{vec}(C)) \\ &\quad + (\sigma_L (B \otimes B) \otimes I_{d^2} + \mathcal{B}\mathcal{R}) \text{vec}(C) \otimes E(\text{vec}(V_\infty)) \\ &\quad + (\sigma_L I_{d^2} \otimes (B \otimes B) + \mathcal{B}\mathcal{R}) E(\text{vec}(V_\infty)) \otimes \text{vec}(C)]. \end{aligned}$$



## The Increments of $Y$

For  $\Delta > 0$  the sequence of increments  $\mathbf{Y} = (\mathbf{Y}_n)_{n \in \mathbb{N}}$  defined by

$$\mathbf{Y}_n = \int_{(n-1)\Delta}^{n\Delta} \Sigma_{s-}^{1/2} dL_s$$

gives the **log-returns** over consecutive time periods of length  $\Delta$  in a financial context.

### Stationarity:

If  $\Sigma$  (or  $V$ ) is **stationary**, then  **$\mathbf{Y}$  is stationary**.

## Stationary Second Order Structure of the (“Squared”) Increments

Assume that the previous assumptions regarding the second order behaviour are satisfied and  $E(L_1) = 0$ ,  $\text{var}(L_1) = (\sigma_L + \sigma_W)I_d$  for some  $\sigma_W \in \mathbb{R}^+$ , then:

- $(\mathbf{Y}_n)_{n \in \mathbb{N}}$  has finite fourth moments, mean zero and is uncorrelated.

The increments  $\mathbf{Y}$  are white noise with variance:

$$\text{vec}(\text{var}(\mathbf{Y}_1)) = (\sigma_L + \sigma_W)\Delta\mathcal{A}^{-1}(A \otimes I_d + I_d \otimes A)\text{vec}(C)$$

- but the sequence of “squared” increments  $(\mathbf{Y}_n \mathbf{Y}_n^T)_{n \in \mathbb{N}}$  has non-zero autocorrelations which decrease exponentially (from lag one onwards):

$$\text{acov}_{\mathbf{Y}\mathbf{Y}}(h) = e^{\mathcal{A}\Delta h}\mathcal{A}^{-1}(I_{d^2} - e^{-\mathcal{A}\Delta})(\sigma_L + \sigma_W)\text{cov}(\text{vec}(V_\Delta), \text{vec}(\mathbf{Y}_1 \mathbf{Y}_1^*))$$

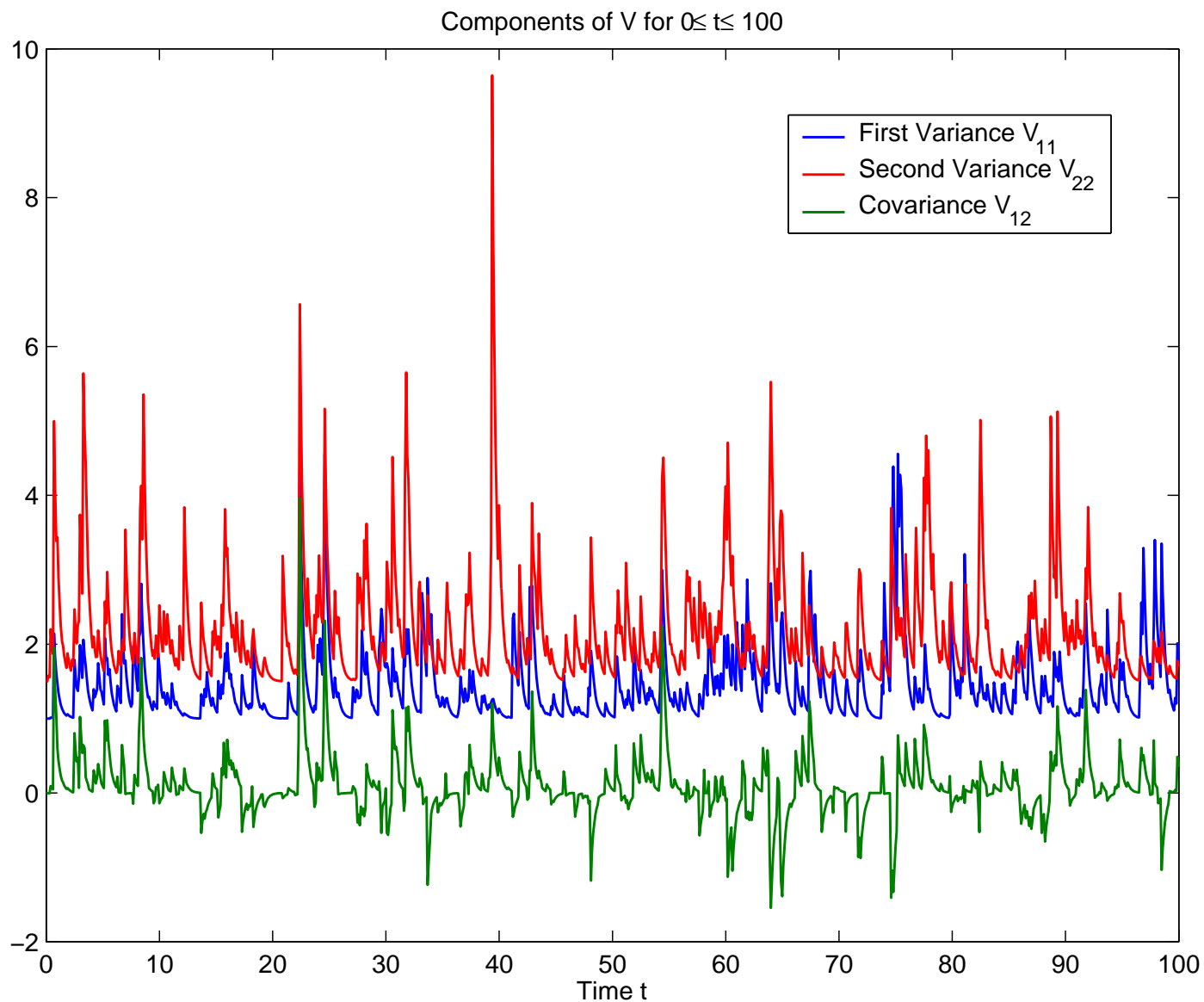
This is the autocovariance structure of an **ARMA(1,1) process**.

## Illustrative Simulations

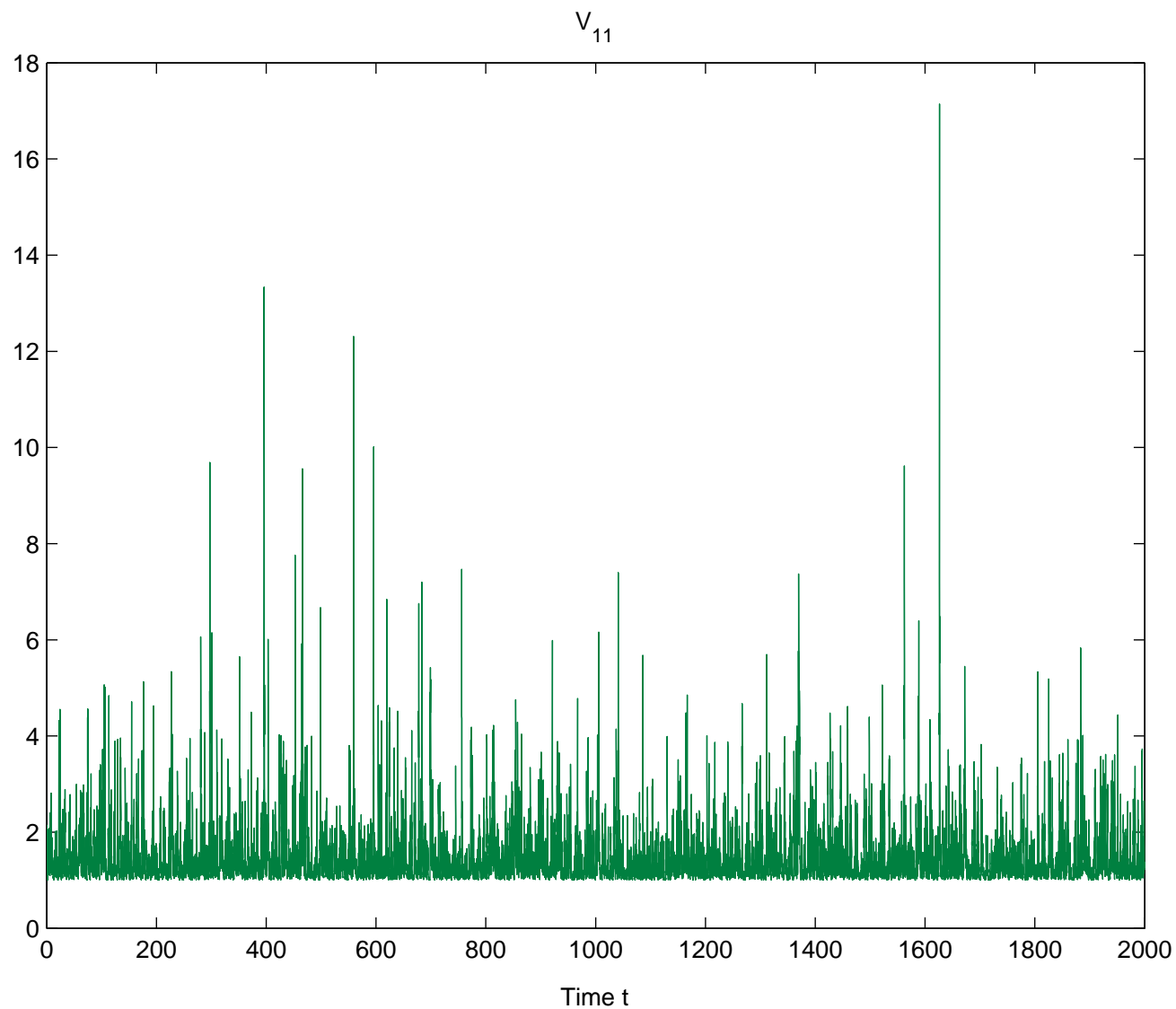
In the following a simulation of a **two-dimensional COGARCH(1,1)** process is shown where:

- the **driving Lévy process** is the sum of a standard **Brownian motion** and a **compound Poisson** process in  $\mathbb{R}^2$  with rate 4 and  $N(0, I_2/4)$ -distributed jumps.
- Hence,  $[L, L]^\diamond$  is a compound Poisson process with Wishart-distributed jumps.
- $A = -1.6I_2$ ,  $B = I_2$  and  
$$C = \begin{pmatrix} 1 & 0 \\ 0 & 1.5 \end{pmatrix}$$
 (corresponds to a “mean” correlation of zero).

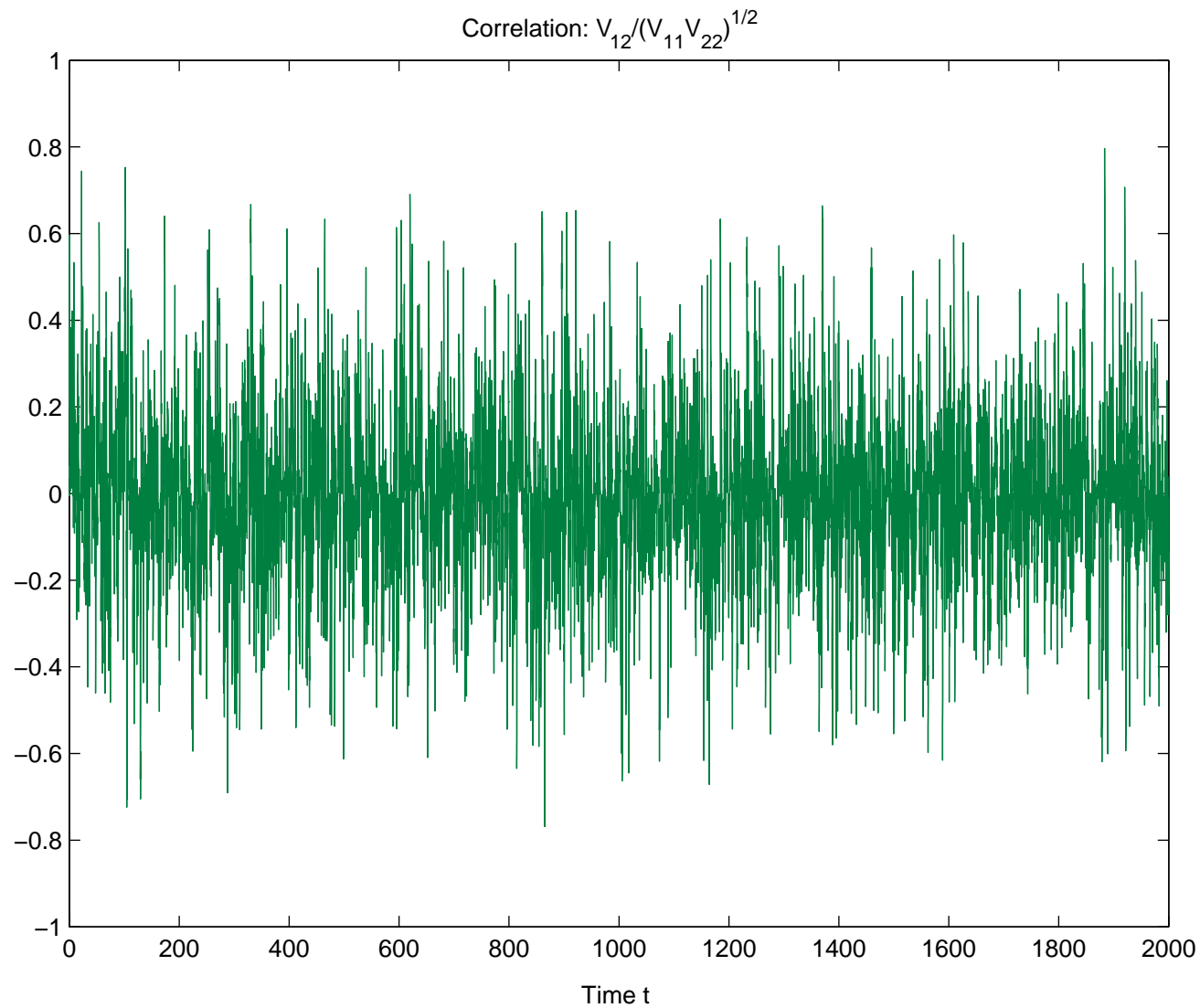
# Stochastic Volatility Process $\Sigma$



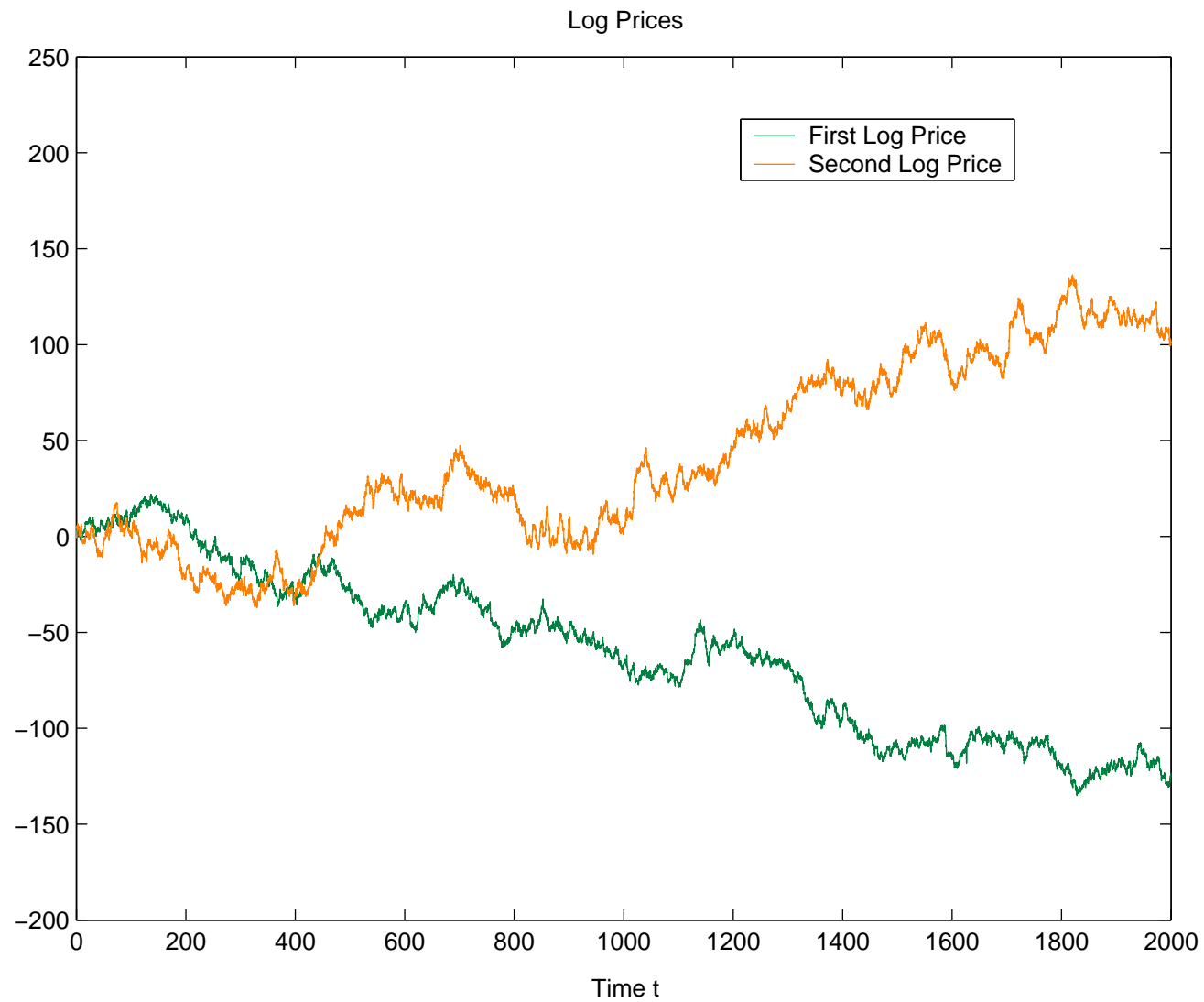
# First Stochastic Variance Process $\Sigma_{11}$



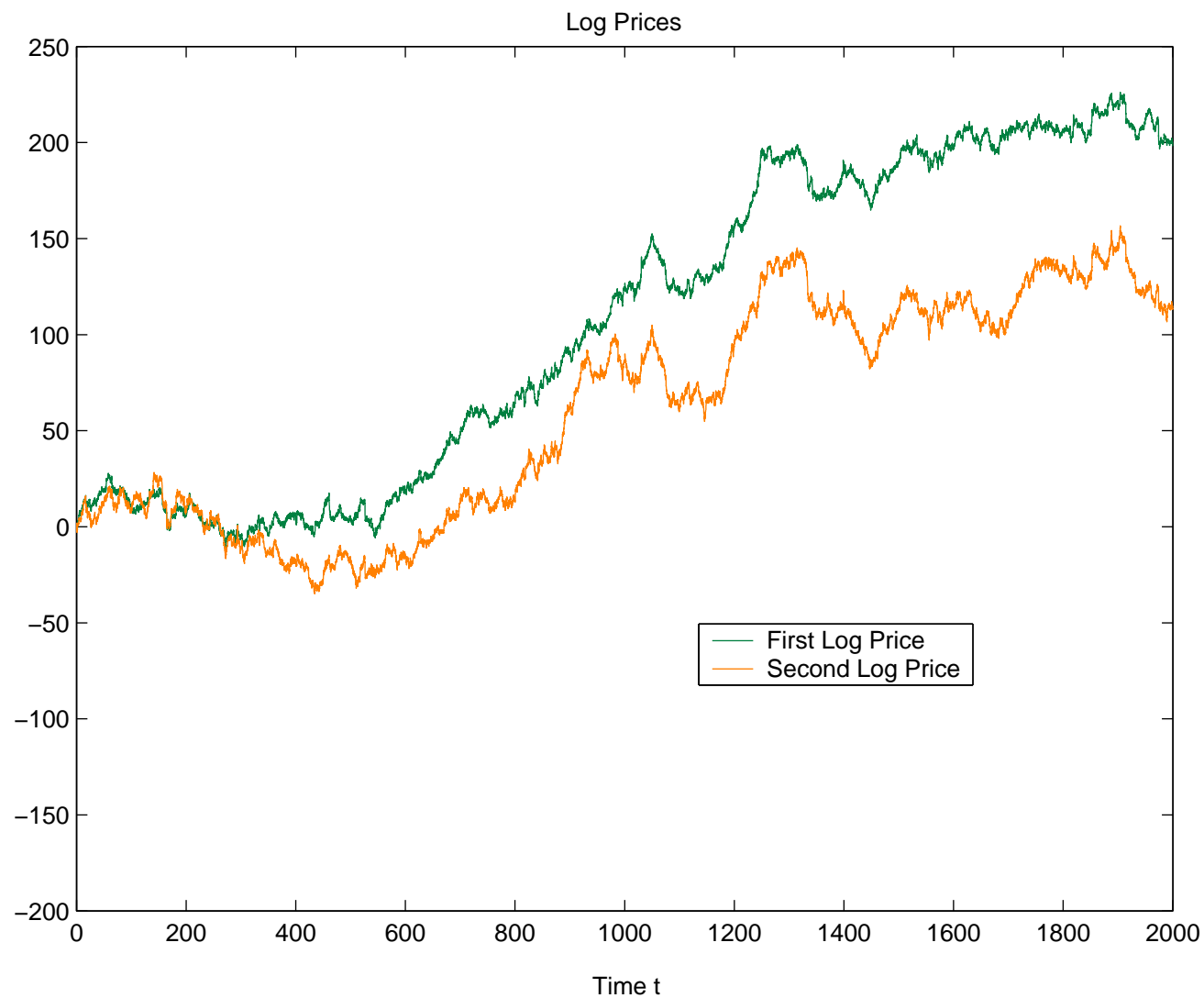
# Stochastic Correlation Process $\Sigma_{12}/\sqrt{\Sigma_{11}\Sigma_{22}}$



# Log-Price Process $Y$

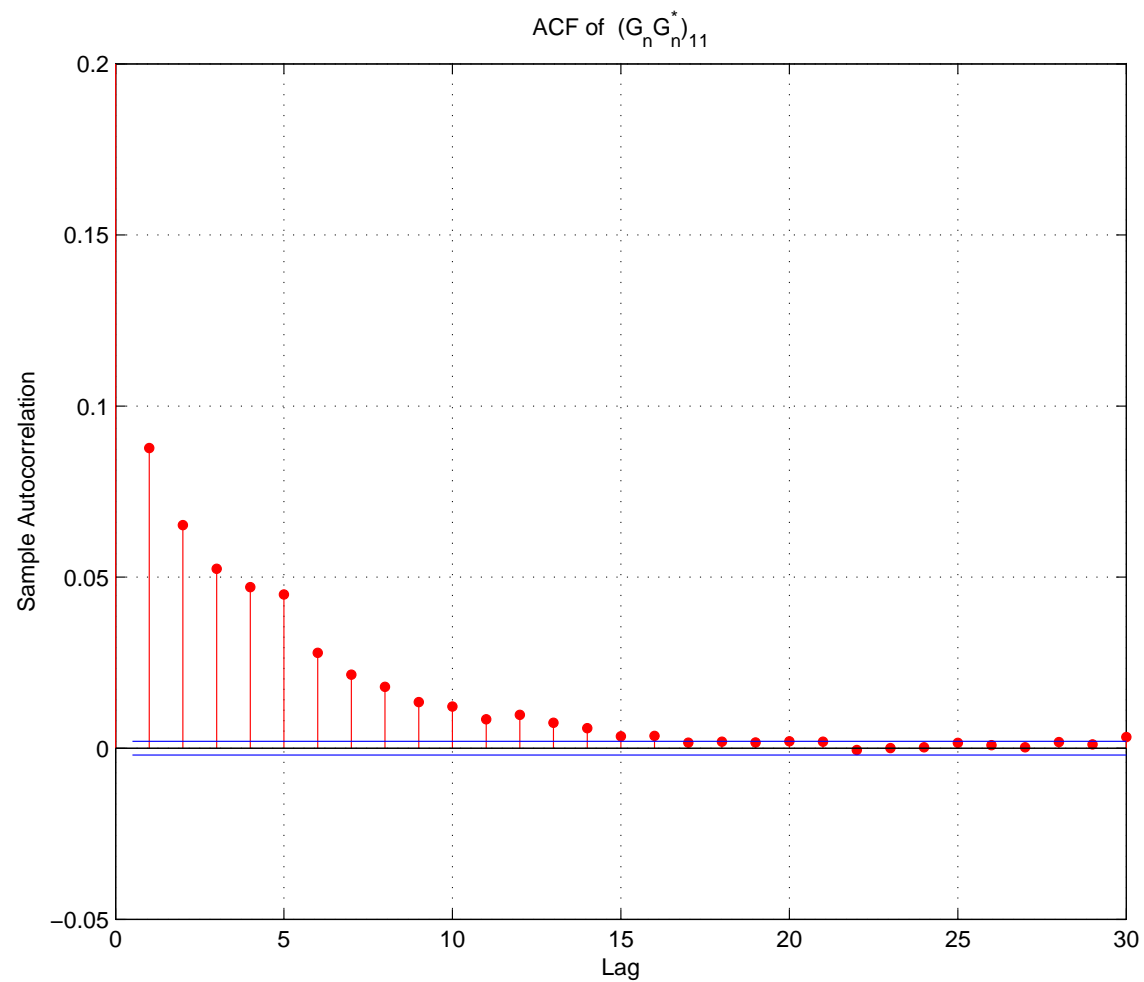


# Alternative $C$ : Log-Price Process $Y$

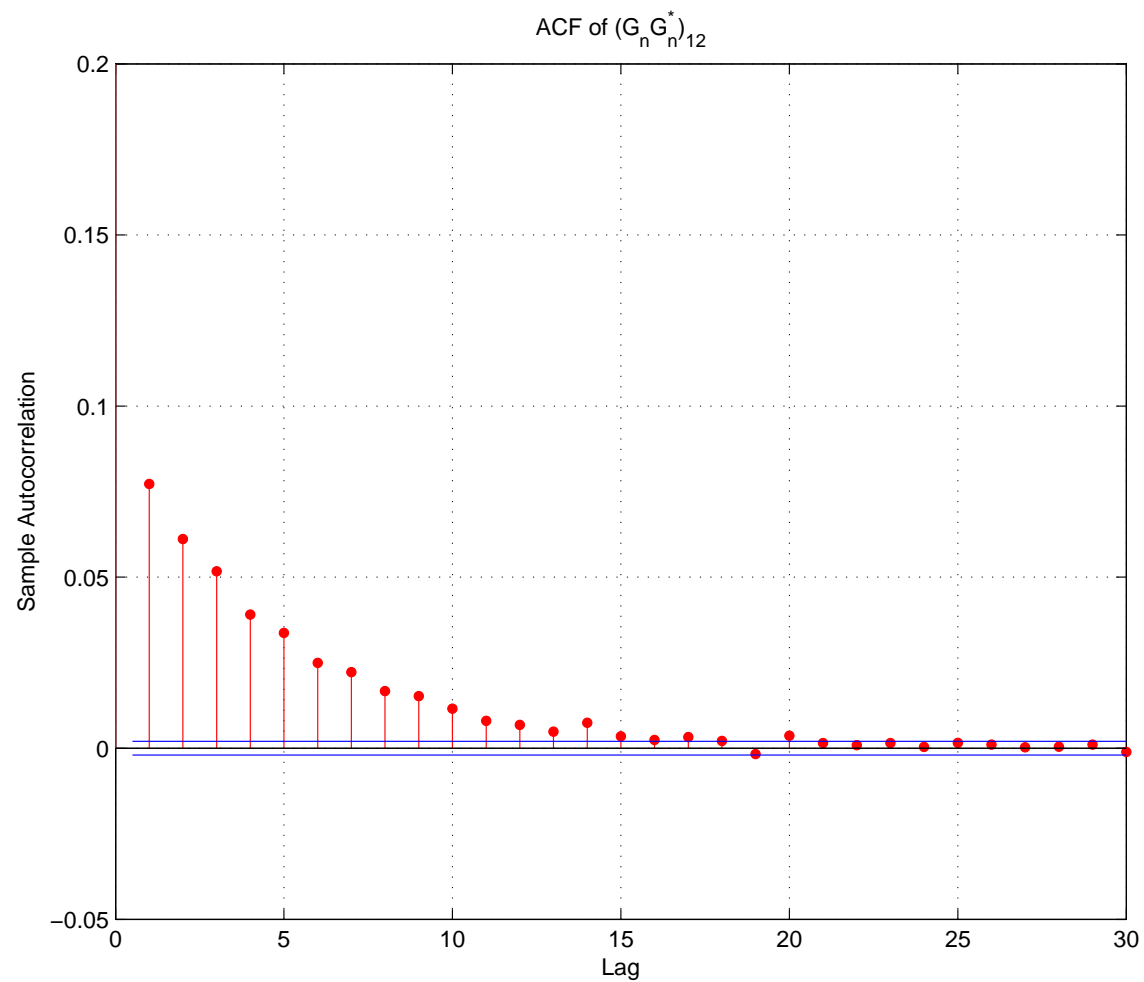




# ACF “Squared Returns” $(YY^T)_{11}$



# ACF “Squared Returns” $(YY^T)_{12}$



**Thank you very much for your attention!**