The Lévy-driven Continuous-Time Garch Model

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Joint work with

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Question: How to model the volatility $(\sigma_t)_{t>0}$.



Figure 1: Deseasonalised 5 minutes log-returns of Intel (February 1 - May 31, 2002) and estimated volatility.

Stylized facts of volatility:

- (1) volatility is random;
- (2) volatility has heavy-tailed marginals (higher moments do not exist);
- (3) volatility has skewed marginals (leverage effect);
- (4) volatility is a stochastic process with long-range dependence effect;
- (5) volatility is a stochastic process with clusters in the extremes.

Recall discrete time GARCH(1,1) model

 $Y_n = \sigma_n Z_n$ i.i.d. innovations $(Z_n)_{n \in \mathbb{N}_0}$,

Volatility process: Define for σ^2 the random recurrence equation

$$\sigma_n^2 = \beta + \lambda Y_{n-1}^2 + \delta \sigma_{n-1}^2, \quad n \in \mathbb{N}.$$

Reorganise and iterate the recurrence:

$$\sigma_n^2 = \beta + \lambda Y_{n-1}^2 + \delta \sigma_{n-1}^2 = \beta + (\delta + \lambda Z_{n-1}^2) \sigma_{n-1}^2$$
$$= \beta \sum_{i=0}^{n-1} \prod_{j=i+1}^{n-1} (\delta + \lambda Z_j^2) + \sigma_0^2 \prod_{j=0}^{n-1} (\delta + \lambda Z_j^2)$$
(1)

Under appropriate conditions: $\sigma_n^2 \xrightarrow{d} \sigma_\infty^2 \stackrel{d}{=} \beta \sum_{i=0}^{\infty} \prod_{j=1}^{i} (\delta + \lambda Z_j^2).$

Continuous time GARCH(1,1)

Idea: start with (1) and replace the sum by an integral

$$\Leftrightarrow \quad \sigma_n^2 = \left(\beta \int_0^n \exp\left(-\sum_{j=0}^{[s]} \log(\delta + \lambda Z_j^2) \, ds\right) + \sigma_0^2\right) \exp\left(\sum_{j=0}^{n-1} \log(\delta + \lambda Z_j^2)\right)$$

Replace Z_j by jumps of a Lévy process L and take $\beta, \eta = -\log \delta, \varphi = \lambda/\delta$. Then for a finite r.v. σ_0^2 define the **volatility process**

$$\sigma_t^2 = \left(\beta \int_0^t e^{X_s} ds + \sigma_0^2\right) e^{-X_t} \quad t \ge 0$$

with auxiliary process

$$X_t = t\eta - \sum_{0 < s \le t} \log(1 + \varphi(\Delta L_s)^2) \quad t \ge 0.$$

Recall: $(L_t)_{t\geq 0}$ is Lévy process if $Ee^{isL_t} = e^{t\psi_L(s)}$, $s \in \mathbb{R}$, with

$$\psi_L(s) = i\gamma_L s - \tau_L^2 \frac{s^2}{2} + \int_{\mathbb{R}} (e^{isx} - 1 - isxI_{\{|x| < 1\}}) \Pi_L(dx), \quad s \in \mathbb{R}.$$

 $(\gamma_L, \tau_L, \Pi_L)$ characteristic tripel, Π_L Lévy measure: $\int_{|x| < \varepsilon} x^2 \Pi_L(dx) < \infty$.

Define the **COGARCH(1,1)process** by

$$G_t = \int_{(0,t]} \sigma_{t-} dL_t \quad t \ge 0.$$

(Note: this defines the martingale part of the price process.)



First: Simulated VG driven COGARCH(1,1) process with $\beta = 0.04$, $\eta = 0.053$ and $\varphi = 0.038$; second: differenced COGARCH process $(G_t^{(1)})$; third: volatility process (σ_t) ; last: VG process (L_t) with characteristic function $Ee^{iuL_1} = (1 + u^2/(2C))^{-C}$ and C = 1;

Properties

- G jumps at the same times as L with jump size $\Delta G_t = \sigma_t \Delta L_t$.
- $(X_t)_{t\geq 0}$ is spectrally negative, has drift η , no Gaussian part, Lévy measure

$$\Pi_{X}([0,\infty)) = 0 \quad \Pi_{X}((-\infty, -x]) = \Pi_{L}(\{|y| \ge \sqrt{(e^{x} - 1)\varphi}\}) \text{ for } x > 0.$$
• $d\sigma_{t}^{2} = (\beta - \eta\sigma_{t-}^{2}) dt + \varphi \sigma_{t-}^{2} d[L, L]_{t}^{(d)}$
where $[L, L]_{t}^{(d)} = \sum_{0 < s \le t} (\Delta L_{s})^{2}$ and
$$\sigma_{t}^{2} = \sigma_{0}^{2} + \beta t - \eta \int_{0}^{t} \sigma_{s}^{2} ds + \varphi \sum_{0 < s \le t} \sigma_{s-}^{2} (\Delta L_{s})^{2} \quad t \ge 0.$$
(2)

•
$$\int_{\mathbb{R}} \log \left(1 + \varphi x^2 \right) \Pi_L(dx) < \eta \quad \Longleftrightarrow \quad EX_1 > 0$$
$$\iff \quad \sigma_t^2 \xrightarrow{d} \sigma_\infty^2 \xrightarrow{d} \beta \int_0^\infty e^{-X_t} dt.$$

Sample path behaviour

- From (2) we know that σ_t^2 has only upwards jumps.
- If $(L_t)_{t\geq 0}$ is compound Poisson with jump times $0 = T_0 < T_1 < \ldots$,

$$\sigma_t^2 = \frac{\beta}{\eta} + \left(\sigma_{T_j}^2 - \frac{\beta}{\eta}\right) e^{-(t - T_j)\eta}, \quad t \in (T_j, T_{j+1}).$$

• For the stationary process, we have $\sigma_{\infty}^2 \ge \frac{\beta}{\eta}$ a.s.



Sample paths of σ_t^2 (solid line) and $\hat{\sigma}_t^2$ (dotted line) of one simulation of a VG process.

Theorem Suppose that $EL_1 = 0$, $var(L_1) = 1$. Define $Ee^{-sX_t} = e^{t\Psi_X(s)}$. Assume that the volatility process is stationary, and define $G_i^{(1)} := \int_{i-r}^i \sigma_{s-} dL_s$. If $\Psi_X(1) < 0$, then

$$EG_i^{(1)} = 0, \ E(G_i^{(1)})^2 = \frac{r\beta}{-\Psi_X(1)} EL_1^2 \text{ and } \operatorname{corr}(G_i^{(1)}, G_{i+h}^{(1)}) = 0.$$

If $EL_1^4 < \infty$, $\Psi_X(2) < 0$ and $\int_{\mathbb{R}} x^3 \nu_L(dx) = 0$, then for $k, p > 0$

$$\operatorname{corr}((G_i^{(1)})^2, (G_{i+h}^{(1)})^2) = ke^{-hp}, \quad h \in \mathbb{N}.$$

Theorem Assume that L_1 is symmetric and that there exists $\kappa > 0$ such that $|L_1|^{\kappa} \log^+ |L_1| < \infty$ and $\Psi_X(\kappa/2) = 1$.

Then a stationary version of the volatility process exists with

$$P(\sigma_t > x) \sim cx^{-\kappa/2}, \quad x \to \infty.$$

Stylized facts of volatility:

- (1) volatility is random;
- (2) volatility has heavy-tailed marginals (higher moments do not exist:K., Lindner and Maller (2004), Fasen, K., Lindner (2004));
- (3) volatility has skewed marginals (leverage effect introduced in Haug et al.)
- (4) volatility is a stochastic process with long-range dependence effect (acf decreases geometrically: K., Lindner and Maller (2004));
- (5) volatility is a stochastic process with clusters in the extremes:Fasen: Extremes of genOU processes (2006, 2007).

Question: Can we find a discrete time skeleton, which approximates the COGARCH(1,1) process, and is a GARCH(1,1) process.

The following approximation, called **first jump approximation** shows that (under some technical conditions) the solution of a Lévy-driven SDE can be approximated arbitrarily close, by replacing the Lévy process with its first jump approximation.

Theorem [Szimayer and Maller (2007), Haug and Stelzer (2007)] Let L be a Lévy process in \mathbb{R}^d , which has no Brownian part, drift γ_L and Lévy measure Π_L and satisfies $EL^2(1) = 1$.

For $n \in \mathbb{N}$ let $1 > \varepsilon^{(n)} \downarrow 0$ and $0 = t_0^{(n)} < t_1^{(n)} < t_2^{(n)} \dots \uparrow \infty$. Set $\delta^{(n)} := \sup_{i \in \mathbb{N}} (t_i^{(n)} - t_{i-1}^{(n)})$ and assume that $\lim_{n \to \infty} \delta^{(n)} = 0$. Assume that

$$\lim_{n \to \infty} \delta^{(n)} (\Pi(\{x \in \mathbb{R}^d : |x| > \varepsilon^{(n)}\})^2 = 0.$$
(3)

Define for all $n \in \mathbb{N}$

$$\begin{split} \gamma^{(n)} &:= \gamma_L - \int_{\varepsilon^{(n)} < |x| \le 1} x \Pi_L(dx) \\ \tau_i^{(n)} &:= \inf\{t : t_{i-1}^{(n)} < t \le t_i^{(n)}, \, |\Delta L_t| > \varepsilon^{(n)}\} \quad \forall i \in \mathbb{N} \\ \widetilde{L}_t^{(n)} &:= \gamma^{(n)} t + \sum_{\{i \in \mathbb{N} : \tau_i^{(n)} \le t\}} \Delta L_{\tau_i^{(n)}} \quad \forall t \ge 0 \\ \overline{L}_t^{(n)} &:= \widetilde{L}_{t_{i-1}^{(n)}}^{(n)}. \end{split}$$

Then

$$\widetilde{L}^{(n)} \to L$$
 in ucp as $n \to \infty$ and $d_S(\overline{L}^{(n)}, L) \xrightarrow{P} 0 \quad n \to \infty$.

Remark (i) Whenever one of the sequences $(\delta^{(n)})$ or $(\varepsilon^{(n)})$ are given, one can always choose the other such that (3) holds.

(ii) Note that the time grid is not necessarily equidistant. The construction allows for discrete sampling of a continuous-time Lévy-driven model. This is useful for high-frequency data.

(iii) The construction allows also the embedding of a discrete-time model into a continuous-time jump model.

Example [COGARCH(1,1) and its GARCH(1,1) approximation]

Maller, Müller and Szimayer (2007) specify this approach and apply it to:

- (1) Parameter estimation by pseudo MLE.
- (2) Option pricing using the approach of Ritchken and Trevor (1999).

For an alternative approach, see Kallsen and Vesenmayer (2007).

Example [COGARCH(1,1) and its GARCH(1,1) approximation, Maller, Müller and Szimayer (2007)]

We use the notation as in the theorem and assume that all assumptions hold. For $n \in \mathbb{N}$ set $\Delta t_i(n) := t_i^{(n)} - t_{i-1}^{(n)}$ and define $\Delta L_{\tau_i^{(n)}}$ as the first jump of size larger than $\varepsilon^{(n)}$ in $(t_{i-1}^{(n)}, t_i^{(n)}]$. Define

$$Z_{i,n} = \frac{\mathbf{1}_{\{\tau_i^{(n)}\} < \infty} \Delta L_{\tau_i^{(n)}} - \nu_i^{(n)}}{\xi_i^{(n)}}, \quad i \in \mathbb{N}.$$

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By the strong Markov property $(\mathbf{1}_{\{\tau_i^{(n)}<\infty\}}\Delta L_{\tau_i^{(n)}})_{i\in\mathbb{N}}$ is an iid sequence with distribution

$$\frac{\Pi(dx)\mathbf{1}_{\{|x|>\varepsilon^{(n)}\}}}{\Pi(\{x\in\mathbb{R}^d:|x|>\varepsilon^{(n)}\})}\Big(1-e^{-\eta\Delta t_i(n)\Pi(\{x\in\mathbb{R}^d:|x|>\varepsilon^{(n)}\})}\Big), \quad x\in\mathbb{R}\setminus\{0\}.$$

Then $(Z_{i,n})_{i \in \mathbb{N}}$ is an iid sequence with mean 0 and variance 1. Now recall

$$d\sigma_t^2 = (\beta - \eta \sigma_{t-}^2) \, dt + \varphi \, \sigma_{t-}^2 d[L, L]_t^{(d)} \quad \text{and} \quad G_t = \int_{(0,t]} \sigma_{t-} dL_t \quad t > 0.$$

We discretise as follows: for $G_{0,n} = G_0 = 0$ set

$$G_{i,n} - G_{i-1,n} = \sigma_{i-1,n} \sqrt{\Delta t_i(n)} Z_{i,n}, \quad i \in \mathbb{N},$$

and

$$\sigma_{i,n}^2 = \beta \Delta t_i(n) + \left(1 + \varphi \Delta t_i(n) Z_{i,n}^2\right) e^{-\eta \Delta t_i(n)} \sigma_{i-1,n}^2, \quad i \in \mathbb{N}.$$

This defines a discrete time GARCH(1,1) random recurrence equation; cf. p. 4.

Follow the construction as before and introduce continuous-time versions (piecewise constant) of the auxiliary process $X_{i,n}$, $\sigma_{i,n}^2$ and $G_{i,n}$. Then with the usual technical efforts, it is shown that

$$d_S((G_n, \sigma_n^2), (G, \sigma^2)) \xrightarrow{P} 0 \quad n \to \infty.$$

Question: Can we define a reasonable multivariate COGARCH model.

Definition [Multivariate COGARCH(1,1) model, Stelzer (2007)]

Let L be a d-dimensional Lévy process and $A, B \in M_d(\mathbb{R})$ (the $d \times d$ matrices), $C \in \mathbb{S}_d^+$ (the $d \times d$ positive semi-definite matrices) and set $[L, L]_t^{(d)} := \sum_{0 < s \leq t} \Delta L_s(\Delta L_s)^*$. Then the process $G = (G_t)_{t \in \mathbb{R}^+}$ solving

$$dG_t = V_{t-}^{1/2} dL_t$$

$$dV_t = (B(V_{t-} - C) + (Y_{t-} - C)B^*)dt + AV_{t-}^{1/2} d[L, L]_t^{(d)} V_{t-}^{1/2} A^*$$

with $G_0 \in \mathbb{R}^d$ and $Y_0 \in \mathbb{S}_d^+$ is a **multivariate COGARCH(1,1) process**. **Note:** This definition agrees for d = 1 with the COGARCH(1,1) process. For details see work by Robert Stelzer. **Question:** Can we find a class of models, where the COGARCH(1,1) and the Barndorff-Nielsen and Shephard model belong to.

Recall the COGARCH(1,1) volatility process:

$$\sigma_t^2 = e^{-X_{t-}} \left(\beta \int_0^t e^{X_s} ds + \sigma_0^2 \right) \quad t \ge 0$$

where $(X_t)_{t\geq 0}$ is a spectrally negative Lévy process with positive drift.

Compare to the Barndorff-Nielsen and Shephard OU process

$$\widetilde{\sigma}_t^2 = e^{-\alpha t} \left(\int_0^t e^{\alpha s} dL_{\alpha s} + \widetilde{\sigma}_0^2 \right) \quad t \ge 0$$

where $(L_t)_{t\geq 0}$ is a subordinator.

This motivates the definition of the **generalised Ornstein-Uhlenbeck (genOU) process**

$$V_t = e^{-\xi_t} \left(\int_0^t e^{\xi_{s-}} d\eta_s + V_0 \right) \quad t \ge 0$$
(4)

where $(\xi_t, \eta_t)_{t \ge 0}$ is a bivariate Lévy process and V_0 is an independent starting random variable.

This process has global properties concerning stationarity and second order behaviour, which explains the similarity between the COGARCH model and the BN-S OU model.

The similarity breaks down for the extremal behaviour.

See work by Vicky Fasen and Alexander Lindner.

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